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Quadratic Programming Feature Selection

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Abstract

Identifying a subset of features that preserves classification accuracy is a problem of growing importance, because of the increasing size and dimensionality of real-world data sets. We propose a new feature selection method, named Quadratic Programming Feature Selection (QPFS), that reduces the task to a quadratic optimization problem. In order to limit the computational complexity of solving the optimization problem, QPFS uses the Nyström method for approximate matrix diagonalization. QPFS is thus capable of dealing with very large data sets, for which the use of other methods is computationally expensive. In experiments with small and medium data sets, the QPFS method leads to classification accuracy similar to that of other successful techniques. For large data sets, QPFS is superior in terms of computational efficiency.

Keywords: feature selection, quadratic programming, Nyström method, large data set, high-dimensional data

1. Introduction

The task of feature selection is to reduce the number of variables used in training a classifier. Three main benefits can be drawn from successful feature selection: first, a substantial gain in computational efficiency (especially important for any application that requires classifier execution in real-time); second, scientific discovery by determining which features are most correlated with the class labels (which may in turn reveal unknown relationships among features); and, third, reduction of the risk of overfitting if too few training instances are available (a serious problem particularly in situations with high dimensionalities relative to training set sizes). Document categorization (Forman, 2008), prosthesis control (Momen et al., 2007; Shenoy et al., 2008), cardiac arrhythmia classifica-

tion (Rodriguez et al., 2005), fMRI analysis, gene selection from microarray data (Ding and Peng, 2005; Li et al., 2004; Zhang et al., 2008), real-time identification of polymers (Leitner et al., 2003), and credit card fraud detection are some real-life domains where these gains are especially meaningful.

Many methods have been suggested to solve the variable selection problem. They can be categorized into three groups. *Filter* methods perform feature selection that is independent of the classifier (Bekkerman et al., 2003; Forman, 2003, 2008). *Wrapper* methods use search techniques to select candidate subsets of variables and evaluate their fitness based on classification accuracy (John et al., 1994; Kohavi and John, 1997; Langley, 1994). Finally, *embedded* methods incorporate feature selection in the classifier objective function or algorithm (Breiman et al., 1984; Weston et al., 2001).

Filter methods are often preferable to other selection methods because of their usability with alternative classifiers, their computational speed, and their simplicity (Guyon, 2003; Yu and Liu, 2003). But filter algorithms often score variables separately from each other, so they do not achieve the goal of finding combinations of variables that give the best classification performance. It has been shown that simply combining good variables does not necessary lead to good classification accuracy (Cover, 1974; Cover and Thomas, 1991; Jain et al., 2000). Therefore, one common improvement direction for filter algorithms is to consider dependencies among variables. In this direction approaches based on mutual information, in particular Maximal Dependency (MaxDep) and minimal-Redundancy-Maximum-Relevance (mRMR), have been significant advances (Peng et al., 2005).

The central idea of the MaxDep approach is to find a subset of features which jointly have the largest dependency on the target class. However, it is often infeasible to compute the joint density functions of all features and of all features with the class. One approach to making the MaxDep approach practical is Maximal Relevance (MaxRel) feature selection (Peng et al., 2005). This approach selects those features that have highest relevance (mutual information) to the target class. The main limitation of MaxRel is not accounting for redundancy among features. The mRMR criterion is another version of MaxDep that chooses a subset of features with both minimum redundancy (approximated as the mean value of the mutual information between each pair of variables in the subset) and maximum relevance (estimated as the mean value of the mutual information between each feature and the target class). Given the prohibitive cost of considering all possible subsets of features, the mRMR algorithm selects features greedily, minimizing their redundancy with features chosen in previous steps and maximizing their relevance to the class.

The new method proposed in this paper aims at dealing with very large data sets with high dimensionality providing a time complexity improvement respect to current methods.We show how to build on well-established mathematical methods to reduce time and space complexity. The new method is named Quadratic Programming Feature Selection (QPFS) because it is based on efficient quadratic programming (Bertsekas, 1999). We introduce an objective function with quadratic and linear terms. The quadratic term captures the dependence (that is, similarity, correlation, or mutual information) between each pair of variables, whereas the linear term captures the relationship between each feature and the class label. For large data sets, solving a quadratic programming problem can have high time and space complexity. Therefore, we show how to reformulate the optimization problem in a lower dimensional subspace using the Nyström method for matrix diagonalization (Fowlkes et al., 2001). The Nyström approximation allows the variables to be sampled, without losing much information but with a great improvement in the speed of the algorithm. Ex-

perimental results show that the QPFS method achieves accuracy similar to that of other methods on medium-size data sets, while on the well-known large MNIST data set, QPFS is more efficient than its predecessors.

The present manuscript is organized as follows. Section 2 presents the QPFS algorithm, including the Nyström approximation, error estimation, theoretical complexity, and implementation issues. Section 3 provides a description of data sets, and experimental results in terms of classification accuracy and running time.

2. The QPFS Algorithm

Our goal is to develop a feature selection method capable of succeeding with very large data sets. To achieve this goal, our first contribution is a novel formulation of the task. The new formulation uses quadratic programming, a methodology that has previously been successful for a broad range of other quite different applications (Bertsekas, 1999). Assume the classifier learning problem involves N training samples and M variables (also called attributes or features). A quadratic programming problem is to minimize a multivariate quadratic function subject to linear constraints as follows:

$$\min_x \left\{ \frac{1}{2} x^T Q x - F^T x \right\} . \tag{1}$$

Above, x is an M-dimensional vector, $Q \in \mathbb{R}^{M \times M}$ is a symmetric positive semidefinite matrix, and F is a vector in $\mathbb{R}^M$ with non-negative entries. Applied to the feature selection task, Q represents the similarity among variables (redundancy), and F measures how correlated each feature is with the target class (relevance).

After the quadratic programming optimization problem has been solved, the components of x represent the weight of each feature. Features with higher weights are better variables to use for subsequent classifier training. Since x_i represents the weight of each variable, it is reasonable to enforce the following constraints:

$$\begin{aligned} x_i &\geqslant 0 \text{ for all } i = 1, \ldots, M \\ \sum_{i=1}^{M} x_i &= 1 . \end{aligned}$$

Depending on the learning problem, the quadratic and linear terms can have different relative purposes in the objective function. Therefore, we introduce a scalar parameter α as follows:

$$\min_x \left\{ \frac{1}{2} (1-\alpha) x^T Q x - \alpha F^T x \right\} \tag{2}$$

where x, Q and F are defined as before and $\alpha \in [0,1]$. If $\alpha = 1$, only relevance is considered; the quadratic programming problem becomes linear and equivalent to the MaxRel criterion. On the contrary, if $\alpha = 0$, then only independence between features is considered that is, features with higher weights are those which have lower similarity coefficients with the rest of features. Every data set has its best choice of α to extract the minimum number of features for classification purposes. Nevertheless, a reasonable choice of α must balance the linear and quadratic terms of Equation 2. Thus, we estimate the mean value $\bar{q}$ of the elements of the matrix Q and on the mean value $\bar{f}$ of the

elements of the vector F as

$$\bar{q} = \frac{1}{M^2}\sum_{i=1}^{M}\sum_{j=1}^{M} q_{ij} ,$$
$$\bar{f} = \frac{1}{M}\sum_{i=1}^{M} f_i .$$

Since the relevance and redundancy terms in Equation 2 are balanced when $(1-\hat{\alpha})\bar{q} = \hat{\alpha}\bar{f}$, a reasonable initial estimate of α is

$$\hat{\alpha} = \frac{\bar{q}}{\bar{q}+\bar{f}}.$$

The goal of balancing both terms in the QPFS objective function, Equation 2, is to ensure that both redundancy and relevance are taken into account. If features are only slightly redundant, that is, they have low correlation with each other, then the linear term in Equation 1 is dominant: $\bar{f} \gg \bar{q}$. Making α small reduces this dominance. On the other hand, if the features have a high level of redundancy relative to relevance ($\bar{q} \gg \bar{f}$), then the quadratic term in Equation 1 can dominate the linear one. In this case, overweighting the linear term (α close to 1) makes the objective function be balanced.

Experimental results in Section 3 show that using $\hat{\alpha}$ leads to good results. Alternatively, it is possible to use a validation subset to determine an appropriate value for α. However, that approach requires evaluating the accuracy of the underlying classifier for each point in a grid of α values. In this case, QPFS would become a wrapper feature selection method instead of a filter method because it would need the classifier accuracy to determine the proper value of α.

2.1 Similarity Measures

One advantage of the problem formulation above is that it is sufficiently general to permit any symmetric similarity measure to be used. In the remainder of this paper, the Pearson correlation coefficient and mutual information are chosen, because they are conventional and because they are representative ways to measure similarity.

The Pearson correlation coefficient is simple and has been shown to be effective in a wide variety of feature selection methods, including correlation based feature selection (CFS) (Hall, 2000) and principal component analysis (PCA) (Duda et al., 2000). Formally, the Pearson correlation coefficient is defined as

$$\rho_{ij} = \frac{cov(v_i, v_j)}{\sqrt{var(v_i)var(v_j)}}$$

where *cov* is the covariance of variables and *var* is the variance of each variable. The sample correlation is calculated as

$$\hat{\rho}_{ij} = \frac{\sum_{k=1}^{N}(v_{ki}-\bar{v_i})(v_{kj}-\bar{v_j})}{\sqrt{\sum_{k=1}^{N}(v_{ki}-\bar{v_i})^2\sum_{k=1}^{N}(v_{kj}-\bar{v_j})^2}} \tag{3}$$

where N is the number of samples, v_{ki} is the k-th sample of random variable v_i, and $\bar{v_i}$ is the sample mean of the random variable v_i.

Each matrix element q_{ij} is defined to be the absolute value of the Pearson correlation coefficient of the pair of variables v_i and v_j, that is, $q_{ij} = |\hat{\rho}_{ij}|$. Suppose a classifier learning problem with C classes, the relevance weight of variable v_i, F_i, is computed using a modified correlation coefficient (Hall, 2000) which is an extension to the C-class classification scenario. The modified definition is

$$F_i = \sum_{k=1}^{C} \hat{p}(K=k)|\hat{\rho}_{iC_k}|$$

where K is the target class variable, C_k is a binary variable taking the value 1 when $K = k$ and 0 otherwise, $\hat{p}(K=k)$ is the empirical prior probability of class k, and $\hat{\rho}_{iC_k}$ is the correlation between feature v_i and binary variable C_k, computed according to Equation 3.

Because the correlation coefficient only measures the *linear* relationship between two random variables, it may not be suitable for some classification problems. Mutual information can capture nonlinear dependencies between variables. Formally, the mutual information between two random variables v_i and v_j is defined as

$$I(v_i;v_j) = \int\int p(v_i,v_j)\log\frac{p(v_i,v_j)}{p(v_i)p(v_j)}dv_i dv_j \ .$$

Computing mutual information is based on estimating the probability distributions $p(v_i)$, $p(v_j)$ and $p(v_i,v_j)$. These distributions can be either discretized or estimated by density function methods (Duda et al., 2000). When mutual information is used, the quadratic term is $q_{ij} = I(v_i,v_j)$ and the linear one is $F_i = I(v_i,c)$.

QPFS using mutual information as its similarity measure resembles mRMR, but there is an important difference. The mRMR method selects features greedily, as a function of features chosen in previous steps. In contrast, QPFS is not greedy and provides a ranking of features that takes into account simultaneously the mutual information between all pairs of features and the relevance of each feature to the class label.

2.2 Approximate Solution of the Quadratic Programming Problem

In high-dimensional domains, it is likely that the feature space is redundant. If so, the symmetric matrix Q is singular. We show now how Equation 2 can then be simplified and solved in a space of dimension less than M, thus reducing the computational cost.

Given the diagonalization $Q = U\Lambda U^T$ in decreasing order of eigenvalues, Equation 2 is equivalent to

$$\min_x \left\{ \frac{1}{2}(1-\alpha)x^T U\Lambda U^T x - \alpha F^T x \right\} \ . \tag{4}$$

If the rank of Q is $k \ll M$, then the diagonalization $Q = U\Lambda U^T$ can be written as $Q = \bar{U}\bar{\Lambda}\bar{U}^T$, where $\bar{\Lambda}$ is a diagonal square matrix consisting of the highest k eigenvalues of Q in decreasing order and $\bar{U}$ is a $M \times k$ matrix consisting of the first k eigenvectors of Q. Then, Equation 4 can be rewritten as

$$\min_x \left\{ \frac{1}{2}(1-\alpha)x^T \bar{U}\bar{\Lambda}\bar{U}^T x - \alpha F^T x \right\} \ .$$

Let $y = \bar{U}^T x$ be a vector in $\mathbb{R}^k$. The optimization problem is reduced to minimizing a derived quadratic function in a k-dimensional space:

$$\min_y \left\{ \frac{1}{2}(1-\alpha) y^T \bar{\Lambda} y - \alpha F^T \bar{U} y \right\}$$

under $M+1$ constraints:

$$\begin{aligned} \bar{U} y &\geq \overrightarrow{0} \\ \sum_{i=1}^{M} \sum_{j=1}^{k} \bar{u}_{ij} y_j &= 1. \end{aligned}$$

We can approximate the original vector x as $x \approx \bar{U} y$.

The matrix Q is seldom precisely singular for real world data sets. However, Q can normally be reasonably approximated by a low-rank matrix formed from its $\tilde{k}$ eigenvectors whose eigenvalues are greater than a fixed threshold $\delta > 0$ (Fine et al., 2001). More precisely, let $\tilde{Q} = U \Gamma U^T$ be the $\tilde{k}$-rank approximation of Q, where $\Gamma \in \mathbb{R}^{M \times M}$ is a diagonal matrix consisting of the $\tilde{k}$ highest eigenvalues of Q and the rest of diagonal entries are zero. Then, the approximate quadratic programming problem is formulated as

$$\min_x \left\{ \frac{1}{2}(1-\alpha) x^T U \Gamma U^T x - \alpha F^T x \right\} .$$

Equivalently,

$$\min_y \left\{ \frac{1}{2}(1-\alpha) y^T \widetilde{\Gamma} y - \alpha F^T \widetilde{U} y \right\} \tag{5}$$

where $y = \widetilde{U}^T x \in \mathbb{R}^{\tilde{k}}$, $\widetilde{\Gamma} \in \mathbb{R}^{\tilde{k} \times \tilde{k}}$ is a diagonal matrix with the nonzero eigenvalues of Γ and $\widetilde{U} \in \mathbb{R}^{M \times \tilde{k}}$ the first $\tilde{k}$ eigenvectors of U. The $M+1$ constraints of the optimization problem are defined as

$$\begin{aligned} \widetilde{U} y &\geq \overrightarrow{0} \\ \sum_{i=1}^{M} \sum_{j=1}^{\tilde{k}} \tilde{u}_{ij} y_j &= 1 . \end{aligned}$$

Given the solutions x^* of Equation 2 and $\tilde{x}^*$ of Equation 5, the error of the approximation can be estimated using the following theorem.

Theorem 1 *(Fine et al., 2001) Given $\tilde{Q}$ a $\tilde{k}$-rank approximation of Q, if $(Q - \tilde{Q})$ is positive semidefinite and $tr(Q - \tilde{Q}) \leq \varepsilon$ then the optimal value of the original problem is larger than the optimal objective value of the perturbed problem and their difference is bounded by*

$$\tilde{g}(\tilde{x}^*) \leq g(x^*) \leq \tilde{g}(\tilde{x}^*) + \frac{d^2 l \varepsilon}{2} \tag{6}$$

where l is the number of active constraints in the perturbed problem and d is an upper bound for the coefficients of the original solution.

In our case, $0 \leq x_i \leq 1$ and $d = 1$. The matrix $(Q - \tilde{Q})$ is positive semidefinite since $(Q - \tilde{Q}) = U(\Lambda - \Gamma)U^T$ and $(\Lambda - \Gamma)$ is a diagonal matrix with positive eigenvalues upper bounded by δ. Moreover $\varepsilon \leq (M - \tilde{k})\delta$ and $l \leq M + 1$, so

$$g(x^*) - \tilde{g}(\tilde{x}^*) \leq \frac{l(M-\tilde{k})\delta}{2} \leq \frac{(M+1)(M-\tilde{k})\delta}{2} = \gamma$$

where $g(x)$ and $\tilde{g}(x)$ are defined as

$$g(x) = \frac{1}{2}(1-\alpha)x^T Q x - \alpha F^T x \text{ for } x \in \mathbb{R}^M \tag{7}$$

$$\tilde{g}(x) = \frac{1}{2}(1-\alpha)x^T \tilde{\Gamma} x - \alpha F^T \tilde{U} x \text{ for } x \in \mathbb{R}^{\tilde{k}} . \tag{8}$$

Although the quadratic programming formulation of the feature selection problem is elegant and provides insight, the formulation by itself does not significantly reduce the computational complexity of feature selection. Thus we introduce the idea of applying a Nyström approximation to take advantage of the redundancy that typically makes the matrix Q almost singular. When this is true, the rank of Q is much smaller than M and the Nyström method can approximate eigenvalues and eigenvectors of Q by solving a smaller eigenproblem using only a subset of rows and columns of Q (Fowlkes et al., 2001). Suppose that $k < M$ is the rank of Q which is represented as

$$Q = \begin{pmatrix} A & B \\ B^T & E \end{pmatrix}$$

where $A \in \mathbb{R}^{k \times k}$, $B \in \mathbb{R}^{k \times (M-k)}$, $E \in \mathbb{R}^{(M-k)\times(M-k)}$, and the rows of $[A\ B]$ are independent. Then, the eigenvalues and eigenvectors of Q can be calculated exactly from the submatrix $[A\ B]$ and the diagonalization of A. Let $S = A + A^{-\frac{1}{2}} B B^T A^{-\frac{1}{2}}$ and its diagonalization $S = R\hat{\Sigma}R^T$ then, the highest k eigenvalues of Q are given by $\tilde{\Lambda} = \hat{\Sigma}$ and its associated eigenvectors $\tilde{U}$ are calculated as,

$$\tilde{U} = \begin{pmatrix} A \\ B^T \end{pmatrix} A^{-\frac{1}{2}} R \hat{\Sigma}^{-\frac{1}{2}} .$$

The application of the Nyström method entails some practical issues. First, a prior knowledge of the rank k of Q is, in general, unfeasible and it is necessary to estimate the number of subsamples r to be used in the Nyström approximation. Second, the r rows of $[A\ B]$ should be, ideally, linearly independent. If the rank of Q is greater than r or the rows of $[A\ B]$ are not linearly independent, an approximation of the diagonalization of Q is obtained whose error can be quantified, in general, as $\|E - B^T A^{-1} B\|$. Although the Nyström approximation is not error-free, if the redundancy of the feature space is large enough, then good approximations can be achieved, as shown in the following sections.

When QPFS+Nyström is used, the rule for setting the value of the α parameter is slightly different. In this case, only the the $[A\ B]$ submatrix of Q is known, and it is necessary to use the Nyström approximation $\hat{Q}$ of the original matrix Q,

$$\hat{Q} = (\hat{q_{ij}}) = \begin{pmatrix} A & B \\ B^T & B^T A^{-1} B \end{pmatrix} .$$

Therefore, the mean value of $\hat{Q}$ is computed as

$$\bar{\hat{q}} = \frac{1}{M^2} \sum_{i=1}^{M} \sum_{j=1}^{M} \hat{q_{ij}}.$$

The mean value $\bar{f}$ of the vector F is still calculated using Equation 3, since QPFS+Nyström needs to know all coordinates of F. To sum up, the value of α for the QPFS+Nyström method is

$$\hat{\alpha} = \frac{\bar{\hat{q}}}{\bar{\hat{q}} + \bar{f}}. \tag{9}$$

The algorithm QPFS+Nyström has two levels of approximation.

1. The first level is to approximate the eigenvalues and eigenvectors of the original matrix Q based on only a subset of rows, applying the Nyström method: $\hat{Q} = \hat{U}\hat{\Lambda}\hat{U}^T$. One of the critical issues with the Nyström method is how to choose the subset of rows to use (Fowlkes et al., 2001). Ideally, the number of linearly independent rows of $[A\,B]$ should be the rank of Q. We use uniform sampling without replacement. This technique has been used successfully in other applications (Fowlkes et al., 2001; Williams and Seeger, 2001). Moreover, theoretical performance bounds for the Nyström method with uniform sampling without replacement are known (Kumar et al., 2009). In particular, we use the following theorem.

   Theorem 2 *(Kumar et al., 2009) Let $Q \in \mathbb{R}^{M\times M}$ be a symmetric positive semidefinite Gram (or kernel) matrix. Assume that r columns of Q are sampled uniformly at random without replacement ($r > k$), let $\hat{Q}$ be the rank-k Nyström approximation to Q, and let $\tilde{Q}$ the best rank-k approximation to Q. For $\varepsilon > 0$, if $r \geq \frac{64k}{\varepsilon^4}$, then*

$$E\left[\|Q-\hat{Q}\|_F\right] \leq \|Q-\tilde{Q}\|_F + \varepsilon\left[\left(\frac{M}{r}\sum_{i\in D(r)} Q_{ii}\right)\sqrt{M\sum_{i=1}^{M} Q_{ii}^2}\right]^{\frac{1}{2}} \tag{10}$$

   where $\sum_{i\in D(r)} Q_{ii}$ is the sum of the largest r diagonal entries of Q and $\|\cdot\|_F$ represents the Frobenius norm.

   As $\left(Q-\tilde{Q}\right)$ is a real symmetric positive semidefinite matrix, it is easy to prove that $\|Q-\tilde{Q}\|_F \leq \text{trace}\left(Q-\tilde{Q}\right)$.
   Equation 10 shows that the error in the Nyström approximation decreases with the number of sampled rows, r.

2. The second level of approximation is to solve the quadratic programming problem using the Nyström approximation. As stated in Section 2.2, only eigenvalues higher than a fixed threshold $\delta > 0$ are considered in the rank of matrix $\hat{Q}$. Then, these top $\tilde{k}$ eigenvalues of matrix $\hat{Q}$ are taken to make up a diagonal matrix $\hat{\tilde{\Lambda}} \in \mathbb{R}^{\tilde{k}\times\tilde{k}}$ and let $\hat{\tilde{U}} \in \mathbb{R}^{M\times\tilde{k}}$ the matrix consisting of the eigenvectors associated to $\hat{\tilde{\Lambda}}$. Therefore, the QPFS+Nyström method approximates Q by $\hat{\tilde{Q}} = \hat{\tilde{U}}\hat{\tilde{\Lambda}}\hat{\tilde{U}}^T$ and the quadratic programming problem is defined as,

$$\hat{\tilde{g}}(x) \quad = \quad \frac{1}{2}(1-\alpha)x^T\hat{\tilde{\Lambda}}x - \alpha F^T\hat{\tilde{U}}x \text{ for } x \in \mathbb{R}^{\tilde{k}}$$

and let $\hat{\hat{x}}^*$ be its optimal solution, and $g(x)$ and $\tilde{g}(x)$ be as described in Equations 7 and 8, respectively. The best rank-$\tilde{k}$ approximation to Q is $\tilde{Q} = U\Gamma U^T$ as given in Section 2.2. A bound on the total error in the QPFS+Nyström approximation is obtained following the reasoning in Fine et al. (2001):

$$\begin{aligned} E\left[g(x^*) - \hat{\hat{g}}(\hat{\hat{x}}^*)\right] &\leq E\left[g(\hat{\hat{x}}^*) - \hat{\hat{g}}(\hat{\hat{x}}^*)\right] \\ &\leq \frac{1}{2}(1-\alpha)E\left[(\hat{\hat{x}}^*)^T\left(Q-\hat{\hat{Q}}\right)(\hat{\hat{x}}^*)\right] \\ &\leq \frac{1}{2}E\left[\|Q-\hat{\hat{Q}}\|_2\|\hat{\hat{x}}^*\|_2^2\right] \\ &\leq \frac{1}{2}(M+1)E\left[\|Q-\hat{\hat{Q}}\|_F\right]. \end{aligned}$$

Applying the bound for the Nyström method with uniform sampling without replacement (Equation 10) and the inequality $\|Q-\tilde{Q}\|_F \leq \text{trace}\left(Q-\tilde{Q}\right) \leq (M-\tilde{k})\delta$ yields

$$\begin{aligned} E\left[g(x^*) - \hat{\hat{g}}(\hat{\hat{x}}^*)\right] &\leq \frac{1}{2}(M+1)\left((M-\tilde{k})\delta + \varepsilon\left[\left(\frac{M}{r}\sum_{i\in D(r)} Q_{ii}\right)\sqrt{M\sum_{i=1}^{M} Q_{ii}^2}\right]^{\frac{1}{2}}\right) \\ &\leq \gamma + \frac{\varepsilon}{2}(M+1)\left[\left(\frac{M}{r}\sum_{i\in D(r)} Q_{ii}\right)\sqrt{M\sum_{i=1}^{M} Q_{ii}^2}\right]^{\frac{1}{2}}. \end{aligned}$$

The total error is the sum of the error γ obtained from the approximation of the quadratic programming problem in a subspace (Equation 6) and the error due to the Nyström method.

2.3 Summary of the QPFS+Nyström Method

Figure 1 shows a diagram of the proposed feature selection method, which can be summarized as follows:

1. Compute the F vector representing the dependence of each variable with the class.
2. Choose r rows of Q according to some criterion (typically, uniform sampling without replacement). Arrange the Q matrix so that these r rows are the first ones. Define the $[A\,B]$ matrix to be the first r rows of Q.
3. Set the value of the α parameter according to Equation 9.
4. Apply the Nyström method knowing $[A\,B]$. Obtain an approximation of the eigenvalues and eigenvectors of Q, $\hat{Q} = \hat{U}\hat{\Lambda}\hat{U}^T$.
5. Formulate the quadratic programming (QP) problem in the lower dimensional space $\hat{\hat{Q}} = \hat{\hat{U}}\hat{\hat{\Lambda}}\hat{\hat{U}}^T$.
6. Solve the QP in the subspace to obtain the solution vector y.
7. Return to the original space via $x = \hat{\hat{U}}y$.
8. Rank the variables according to the coefficients of vector x. In case of equal coefficients, rank them by decreasing relevance F_k to the class.

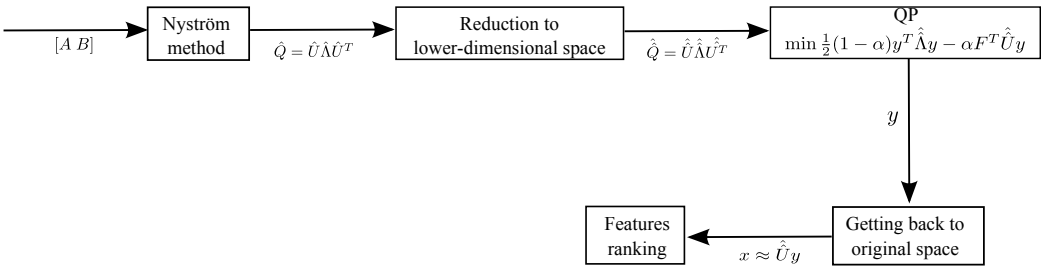

Figure 1: Diagram of the QPFS algorithm using the Nyström method. $[A\ B]$ is the upper $r \times M$ submatrix of Q.

2.4 Complexity Analysis

As already mentioned, the mRMR method is one of the most successful previous methods for feature selection. The main advantage of the QPFS+Nyström method versus mRMR is the time complexity reduction. The time complexities of mRMR and QPFS both have two components, the time needed to compute the matrices Q and F *(Similarity)*, and the time needed to perform variable ranking *(Rank)*. The computational cost of evaluating correlations or mutual information for all variable pairs is $O(NM^2)$ for both mRMR and QPFS. Table 1 shows the time complexities of the three algorithms mRMR, QPFS and QPFS+Nyström.

	mRMR		**QPFS**		**QPFS+Nyström**	
	Similarity	Rank	Similarity	Rank	Similarity	Rank
M large $N \ll pM$	$\mathbf{O(NM^2)}$	$O(M^2)$	$O(NM^2)$	$\mathbf{O(M^3)}$	$O(NpM^2)$	$\mathbf{O(p^2M^3)}$
M medium $N \gg pM$	$\mathbf{O(NM^2)}$	$O(M^2)$	$O(NM^2)$	$\mathbf{O(M^3)}$	$\mathbf{O(NpM^2)}$	$O(p^2M^3)$
M small $N \gg M$	$\mathbf{O(NM^2)}$	$O(M^2)$	$\mathbf{O(NM^2)}$	$O(M^3)$	$\mathbf{O(NpM^2)}$	$O(p^2M^3)$

Table 1: Time complexity of algorithms as a function of training set size N, number of variables M, and Nyström sampling rate p. The predominant cost term is indicated in boldface.

The order-of-magnitude time complexity of QPFS is greater than or similar to that of mRMR. However, the QPFS+Nyström time complexity is lower for $N \gg pM$ and $N \gg M$. When $N \ll pM$, QPFS+Nyström is faster than mRMR if $p^2M^3 \ll NM^2$, that is, when $N \gg p^2M$. For example, if $p = 10^{-2}$ then QPFS+Nyström is more efficient than mRMR if $N \gg 10^{-4}M$, that is, if the size of the training set is greater than 10^{-4} times the number of variables.

2.5 Implementation

We implemented QPFS and mRMR in C using *LAPACK* for matrix operations (Anderson et al., 1990). Quadratic optimization is performed by the Goldfarb and Idnani algorithm implemented in Fortran and used in the R *quadprog* package (Goldfarb and Idnani, 1983; Turlach and Weingessel, 2000).

As mentioned in Section 2.1, in general mutual information computation requires estimating density functions for continuous variables. For simplicity, each variable is discretized in three segments $(-\infty, \mu-\sigma]$, $(\mu-\sigma, \mu+\sigma]$, and $(\mu+\sigma, +\infty)$, where μ is the sample mean of training data and σ its standard deviation. The linear SVM provided by the *LIBSVM* package (Chang and Lin, 2001) was the underlying classifier in all experiments. A linear kernel is used to reduce the number of SVM parameters, thus making meaningful results easier to obtain. Note that mRMR and QPFS can be used with any classifier learning algorithms. We expect results obtained with linear SVMs to be representative.

3. Experiments

The aim of the experiments described here is twofold: first, to compare classification accuracy achieved using mRMR versus QPFS; and second, to compare their computational cost.

3.1 Experimental Design

The data sets used for experiments are shown in Table 2. These data sets were chosen because they are representative of multiple types of classification problems. with respect to the number of samples, the number of features, and the achievable classification accuracy. Moreover, these data sets have been used in other research on the feature selection task (Hua et al., 2009; Lauer et al., 2007; Lecun et al., 1998; Li et al., 2004; Peng et al., 2005; Zhang et al., 2008; Zhu et al., 2008).

In order to estimate classification accuracy, for the ARR, NCI60, SRBCT and GCM data sets 10-fold cross-validation (10CV) and 100 runs were used (Duda et al., 2000). Mean error rates are comparable to the results reported in Li et al. (2004), Peng et al. (2005), Zhang et al. (2008) and Zhu et al. (2008). In the case of the RAT data set, 120 training samples (61 for test) and 300 runs were used, following Hua et al. (2009). The MNIST data set is divided into training and testing subsets as proposed by Chang and Lin (2001), with 60000 and 10000 patterns respectively. Therefore, cross-validation is not done with MNIST.

Time complexity is measured as a function of training set size (N), dimensionality (M), and the Nyström sampling rate ($p = \frac{r}{M}$). In all cases, times are averages over 50 runs. In order to measure time complexity as a function of training set size, the number of SRBCT examples was artificially increased 4 times ($N = 332$) and dimensionality reduced to $M = 140$. Time complexity as a function of dimensionality was measured using the original SRBCT data set, that is, with $N = 83$ and $M = 2308$.

As mentioned above, MaxRel, mRMR and QPFS are all filter methods that can be used with any classifier. Figure 2 shows that the choice of the SVM regularization parameter c does not influence the comparison between mRMR and QPFS. This figure displays the performance of mRMR and QPFS for the ARR data set and different c values. Our goal is not to determine the optimal c value for each data set, but to compare mRMR and QPFS. Therefore, c is set to 1.0 in all experiments.

Data Set	N	M	C	Baseline Error Rate	References
ARR	422	278	2	21.81%	(Peng et al., 2005; Zhang et al., 2008)
NCI60	60	1123	9	38.67%	(Li et al., 2004; Zhang et al., 2008) (Zhu et al., 2008)
SRBCT	83	2308	4	0.22%	(Li et al., 2004; Zhu et al., 2008)
GCM	198	16063	14	33.85%	(Li et al., 2004; Zhang et al., 2008) (Zhu et al., 2008)
RAT	181	8460	2	8.61%	(Hua et al., 2009)
MNIST	60000	780	10	6.02%	(Lauer et al., 2007; Lecun et al., 1998)

Table 2: Description of the data sets used in experiments. N is the number of examples, M is the number of variables, and C is the number of classes. Baseline error rate is the rate obtained taking into account all variables. The last column cites papers where the data sets have been mentioned.

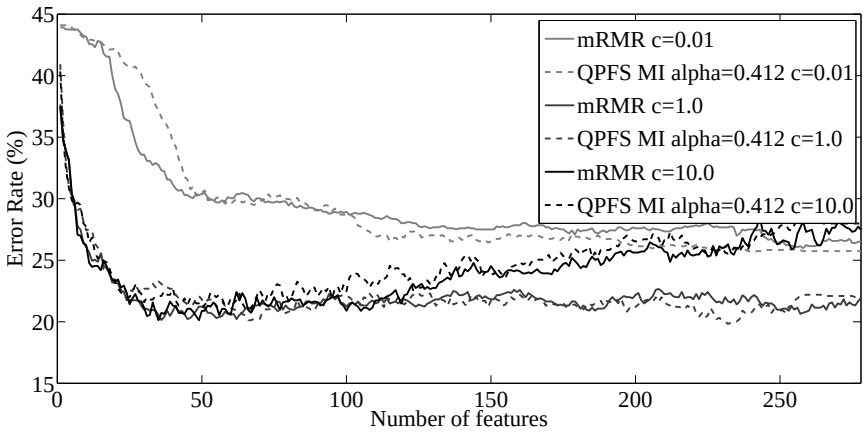

Figure 2: Classification error as a function of the number of features for the ARR data set and different regularization parameter values c in linear SVM. The figure shows that for $c = 0.01$, the SVM is too regularized. The effect when $c = 10.0$ is the opposite and the SVM overfits the training data. A value of $c = 1.0$ is a good tradeoff.

The value of the α parameter chosen for each data set is shown in Table 3. This value is obtained according to Equations 3 and 9. Our hypothesis is that high values of α are better for data sets with high redundancy among variables. On the other hand, if there is low redundancy then small α should yield better results. The Nyström sampling rate p is chosen as large as possible while still yielding a reasonable running time, since larger values reduce error in the approximation of the Q matrix. Other values of the α parameter, $\alpha \in \{0.0, 0.1, 0.3, 0.5, 0.7, 0.9\}$, were considered in all experiments in order to verify that the proposed method of setting α provides good results.

Data Set	p	$\bar{q}$	$\bar{f}$	$\hat{\alpha}$
ARR (cor)	-	0.0889	0.0958	0.481
NCI60 (cor)	-	0.267	0.165	0.618
ARR (MI)	-	0.0106	0.0152	0.411
NCI60 (MI)	-	0.0703	0.253	0.217
SRBCT (MI)	-	0.0188	0.0861	0.179
GCM (MI)	0.05	0.0284	0.158	0.152
RAT (MI)	0.1	0.0346	0.0187	0.649
MNIST (MI)	-	0.0454	0.0515	0.469

Table 3: Values of the α parameter for each data set. Correlation (cor) and mutual information (MI) were used as similarity measures for ARR and NCI60 data sets. Only mutual information was used for SRBCT, GCM, RAT and MNIST data sets. p is the subsampling rate in the Nyström method, $\bar{q}$ is the mean value of the elements of the matrix Q (similarity among each pair of features), and $\bar{f}$ is the mean value of the elements of the F vector (similarity of each feature with the target class). For the MNIST data set only nonzero values have been considered for the statistics due to the high level of sparsity of its features (80.78% sparsity in average).

3.2 Classification Accuracy Results

The aim of the experiments described in this section is to compare classification accuracy achieved with mRMR and with QPFS, with and without Nyström approximation. The MaxRel algorithm (Peng et al., 2005) is also included in the comparison. Two similarity measures, mutual information (MI) and correlation are considered. Classification error is measured as a function of the number of features. We also give results from a baseline method that does random selection of features, in order to determine the absolute advantage of using any feature selection method.

Figure 3 shows the average classification error rate for the ARR data set as a function of the number of features. In Figure 3a, correlation is the similarity measure while mutual information (MI) is applied for Figure 3b. In both cases, the best accuracy is obtained with $\alpha = 0.5$, which means that an equal tradeoff between relevance and redundancy is best. However, accuracies using the values of α specified by our heuristic are similar.

Better accuracy is obtained when MI is used, in which case (Figure 3b) the error rate curve for $\alpha = 0.5$ is similar to that obtained with mRMR. The random selection method yields results significantly worse than those obtained with other algorithms. Comparison with this method shows that the other methods provide a significant benefit up to about 150 features.

For the NCI60 data set (Figure 4), the best accuracy is obtained when mutual information is used (Figure 4b) and α is set to 0.217 according to Table 3. In this case, the accuracy of QPFS is slightly better than the accuracy of mRMR. The value of α close to zero indicates that it is appropriate to give more weight to the quadratic term in QPFS. When correlation is used (Figure 4a), the best accuracy is obtained when α is set according to Equation 3.

Generally, MI as similarity measure leads to better accuracy than correlation. This finding is reasonable given that MI can capture nonlinear relationships between variables. MI is used in the experiments described in the remainder of this section.

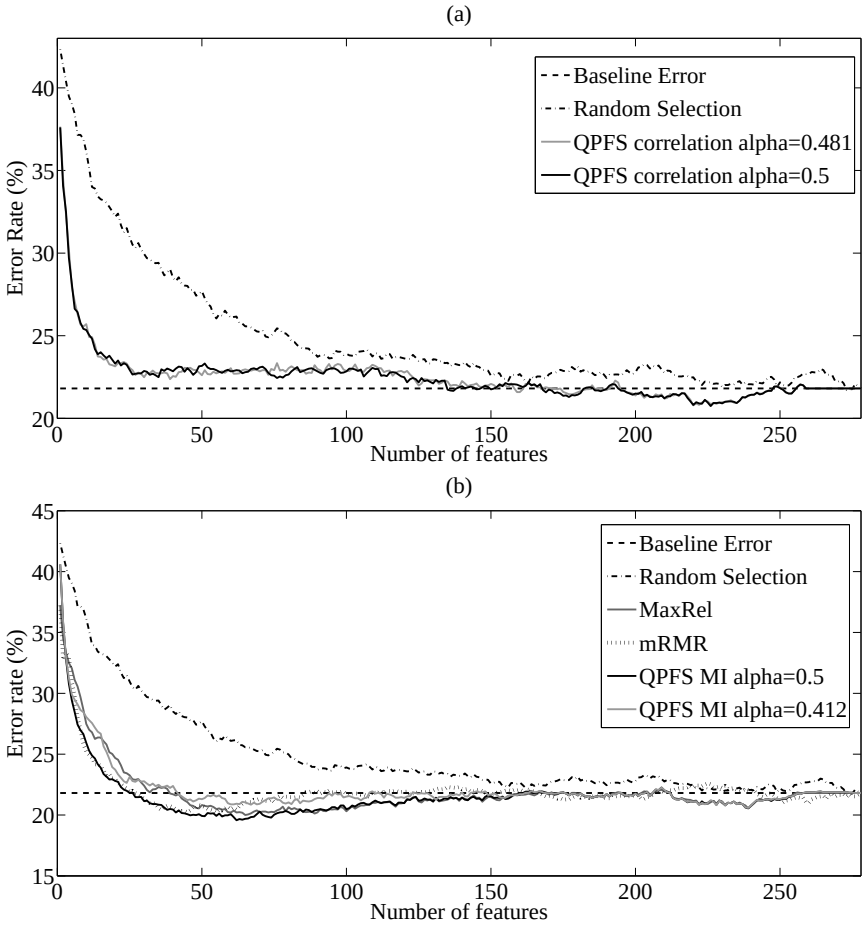

Figure 3: Classification error as a function of the number of features for the ARR data set. **(a)** QPFS results using correlation as similarity measure with different α values. **(b)** MaxRel, mRMR and QPFS results using mutual information as similarity measure and different values of α for QPFS.

Average error rate for the SRBCT data set and different sampling rates as a function of the number of features is shown in Figure 5. Results for the best α value in the grid $\{0, 0.1, 0.3, 0.5, 0.7, 0.9\}$, $\alpha = 0.1$, and the estimated $\hat{\alpha} = 0.179$ are shown in Figure 5a. Accuracies for both α values are similar. The fact that a low value of α is best indicates low redundancy among variables compared to their relevance with the target class. QPFS classification accuracy is similar to that of mRMR. As shown in Figure 5b, when the QPFS+Nyström method is used, the higher the parameter p, the closer the Nyström approximation is to complete diagonalization. QPFS+Nyström gives classification accuracy similar to that of QPFS when $p > 0.1$.

Figure 6 shows error rates for the GCM data set using the algorithms MaxRel, mRMR, and QPFS+Nyström with $\alpha = 0.1$ and $\hat{\alpha} = 0.152$. When the number of features is over 60, accuracy achieved with QPFS+Nyström is better than with mRMR. A sampling rate of 3% is adequate for

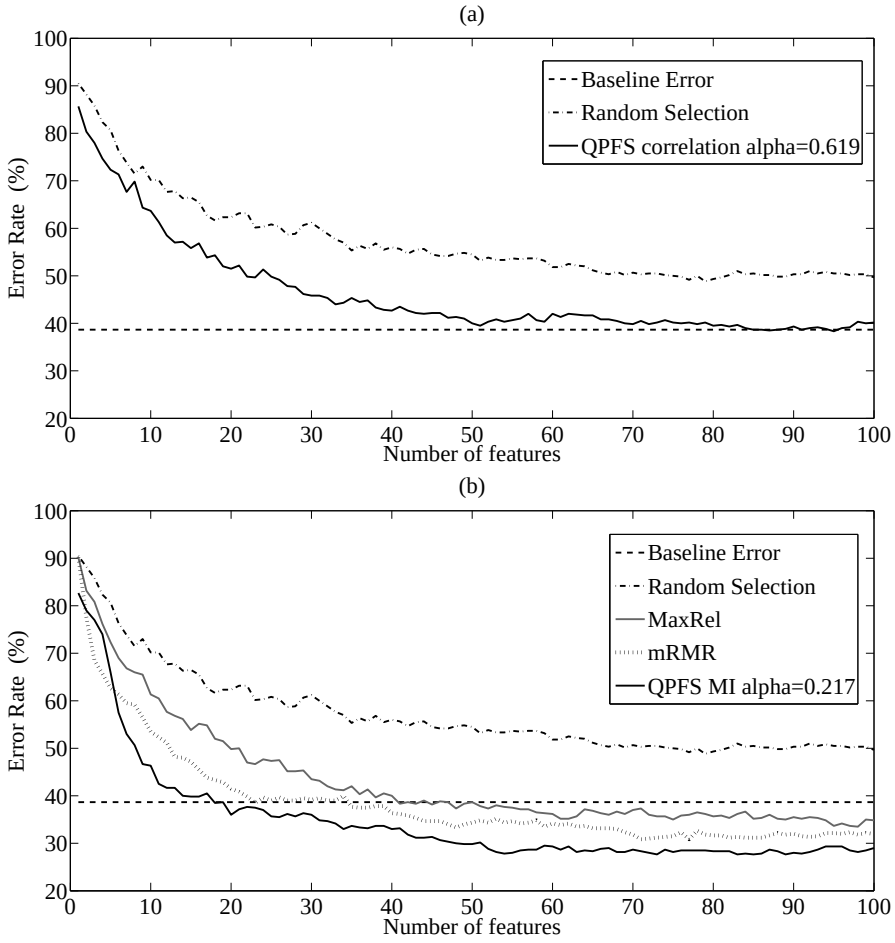

Figure 4: Classification error as a function of number of features for the NCI60 data set. **(a)** QPFS results using correlation as similarity measure with different α values. **(b)** MaxRel, mRMR and QPFS results using mutual information as similarity measure and different values of α for QPFS.

this data set, which represents a major time complexity reduction given a feature space of 16063 variables.

Another data set with many features is the RAT data set, for which Figure 7 shows results. In this case, QPFS+Nyström gives classification accuracy similar to that of mRMR when the subset size is over 80 and the sampling rate is 10%. Given the good performance of the MaxRel algorithm for this data set, it is not surprising that a large α value $\alpha = 0.9$ or $\hat{\alpha} = 0.649$ is best, considering also that QPFS with $\alpha = 1.0$ is equivalent to MaxRel.

The MNIST data set has a high number of training examples. Results for it are shown in Figure 8 for the QPFS with $\alpha = 0.3$, the estimation $\hat{\alpha} = 0.469$ and the QPFS+Nyström with $\hat{\alpha}$ and $p \in \{0.1, 0.2, 0.5\}$. Our C code of mRMR is used instead of the code on the mRMR web site (Peng et al., 2005) which takes a long time to read the training file. The error rate for all algorithms

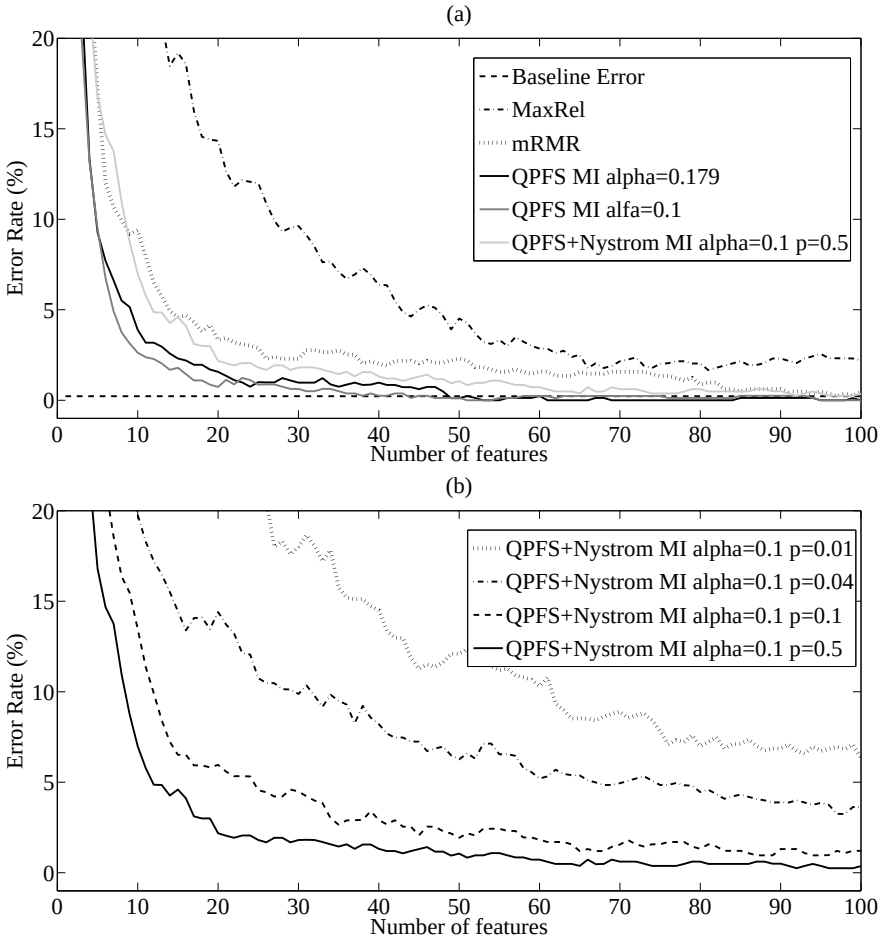

Figure 5: Error rates using MaxRel, mRMR and QPFS+Nyström methods, with mutual information as similarity measure for the SRBCT data set.

reaches a minimum when about 350 features are selected. This is not a surprising fact: analyzing the sparsity of the MNIST features, approximately 400 of them have a level of sparsity higher than 70%. But if the feature space needs a greater reduction, significant differences appears between the studied methods as shown in Figure 8. mRMR and QPFS with $\hat{\alpha} = 0.469$ have similiar performance and close to the best results obtained by QPFS with $\alpha = 0.3$. For this data set, the number of samples is much greater than the number of features, $N \gg M$, and therefore , the time complexity of mRMR and QPFS is the same ($O(NM^2)$). When QPFS+Nyström is applied with $p = 0.2$, the error rate is competitive and the MNIST provides an example of the ability of QPFS+Nyström to handle large data sets reducing the computational cost of mRMR and QPFS by a factor of 5. Note that the error rates shown for the MNIST data set are obtained using a linear kernel. The radial basis function kernel for SVM classifiers is known to lead to lower error rates for the full MNIST data set, but the choice of kernel is an issue separate from feature selection, which is the focus of this paper.

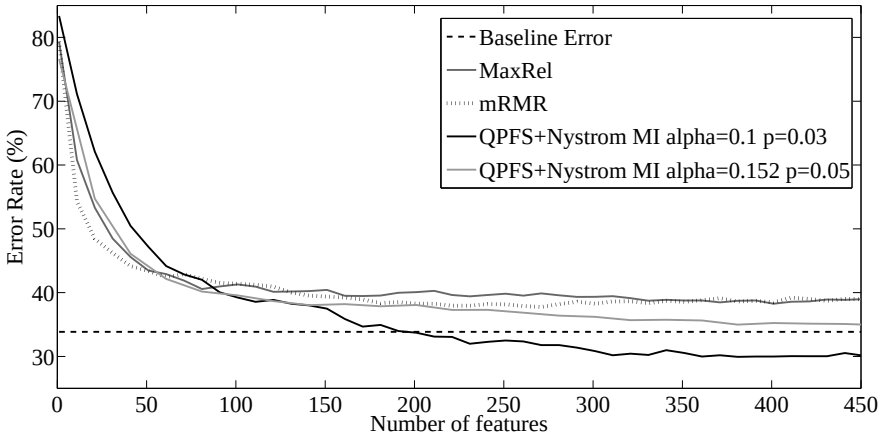

Figure 6: Error rates using MaxRel, mRMR and QPFS+Nyström methods, with mutual information as similarity measure for the GCM data set.

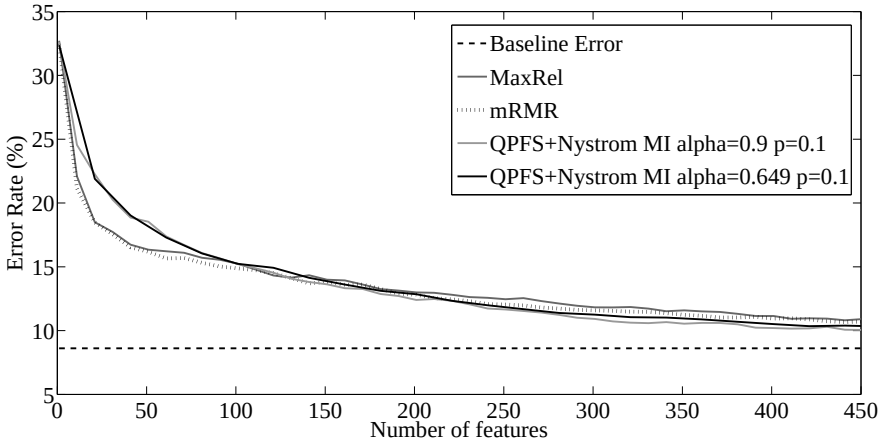

Figure 7: Error rates using MaxRel, mRMR and QPFS+Nyström methods, with mutual information as similarity measure for the RAT data set.

Figure 9 shows a grid of 780 pixels arrayed in the same way as the images in the MNIST data sets. A pixel is black if it corresponds to one of the top 100 (Figure 9a) and 350 (Figure 9b) selected features, and white otherwise. Black pixels are more dense towards the middle of the grid, because that is where the most informative features are. Pixels sometimes appear in a black/white/black checkerboard pattern, because neighboring pixels tend to make each other redundant.

Table 4 evaluates the statistical significance of error rate differences. For each data set, 100 classifiers were trained using the stated number M of selected features. The 100 classifiers arise

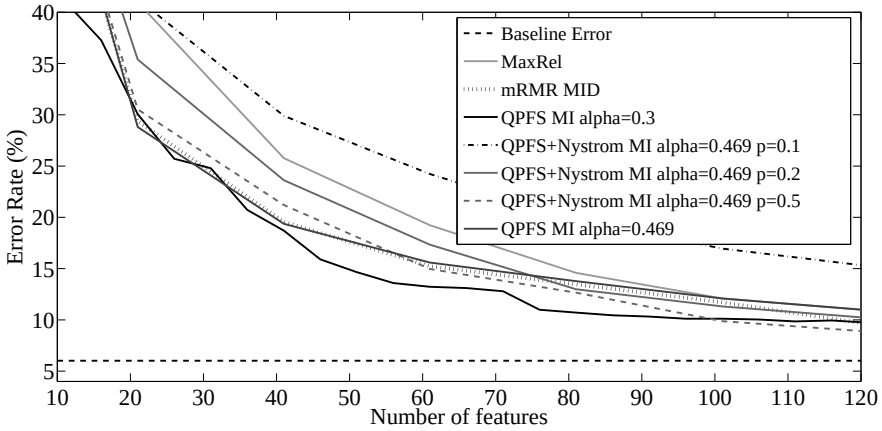

Figure 8: Error rates using MaxRel, mRMR and QPFS+Nyström methods, with mutual information as similarity measure for the MNIST data set.

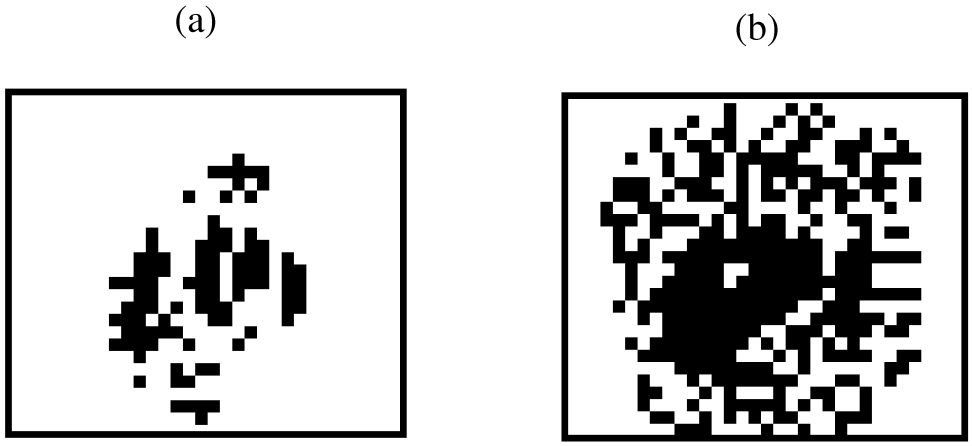

Figure 9: First **(a)** 100 and **(b)** 350 features selected by QPFS+Nyström ($\hat{\alpha} = 0.469$ and $p = 0.5$) for the MNIST data set (black pixels).

from 10 repetitions of 10-fold cross-validation, so which M features are used may be different for each classifier. The one-tailed paired t-test for equal means is applied to the two sets of error rates, one set for mRMR and one set for QPFS. The test is one-tailed because the null hypothesis is that the mRMR method is as good or better than the QPFS method. The test is paired because both methods were applied to the same 100 data set versions. Results of the test are given in the row labeled *significant?*.

For the NCI60 and SRBCT data sets, the best result is obtained when QPFS is used and it is statistically significantly better than mRMR. When 200 to 400 variables are considered, mRMR and QPFS are not statistically significantly different but the accuracy is not as good as in the case of 100 features, probably due to overfitting. In the case of the GCM data set, the mRMR method is statistically significantly better when fewer than 50 variables are considered. If the number of features is over 100, the accuracy with QPFS is significantly better than with mRMR, and the best performance is obtained in this case. For the ARR data set, mRMR is statistically significantly better than QPFS if fewer than 10 features are considered but the error rate obtained can be improved if more features are taken into account. When more than 50 featues are selected, the two methods are not statistically significantly different. The RAT data set behavior is quite similar. When fewer than 100 features are used, the mRMR algorithm is satatistically better than QPFS, but the error rate can be reduced adding more features. The two algorithms are not statistically significantly different in the other cases, except if more than 400 features are involved in which case QPFS is statistically significantly better than mRMR. Note that the error rates shown for QPFS are obtained with the proposed estimation of $\hat{\alpha}$. In some cases, as shown in Figures 3 to 7, this α value is not the best choice.

Beyond simple binary statistical significance, Table 4 indicates that the QPFS method is statistically significantly better when the value of $\hat{\alpha}$ is small. A possible explanation for this finding is the following. When $\hat{\alpha}$ is small, features are highly correlated with the label ($\overline{f} \gg \overline{q}$). The mRMR method is greedy, and only takes into account redundancy among features selected in previous iterations. When features are highly correlated with the label, then mRMR selects features with high relevance and mostly ignores redundancy. In contrast, QPFS evaluates all variables simultaneously, and always balances relevance and redundancy.

3.2.1 COMPARISON WITH OTHER FEATURE SELECTION METHODS

The experiments of this work are focused in comparing QPFS with the greedy filter-type method mRMR (difference form, named MID) which also takes into account the difference between redundancy and relevance. Nevertheless, other feature selection methods independent of the classifier have been considered in the described experiments:

- **mRMR (quotient form, named MIQ)** (Ding and Peng, 2005). While in mRMR (MID form) the difference between the estimation of redundancy and relevance is considered, in the case of mRMR (MIQ form) the quotient of both approximations is calculated.

- **reliefF** (Robnik-Šikonja and Kononenko, 2003). The main idea of ReliefF is to evaluate the quality of a feature according to how well it distinguishes between instances that are near to each other. This algorithm is efficient in problems with strong dependencies between attributes.

- **Streamwise Feature Selection (SFS)** (Zhou et al., 2006). SFS selects a feature if the *benefit* of adding it to the model is greater than the increase in the model complexity. The algorithm scales well to large feature sets and considers features sequentially for addition to a model making unnecessary to know all the features in advance.

Average error rates for MaxRel, mRMR (MID), mRMR (MIQ), reliefF and QPFS using linear SVM ($c = 1.0$) and different number of features are shown in Table 5. Table 6 shows the error rate

	M				
	10	50	100	200	400
RAT					
mRMR	21.15 ± 0.31	16.18 ± 0.27	14.88 ± 0.24	12.81 ± 0.25	10.95 ± 0.23
QPFS $\hat{\alpha} = 0.65$	27.13 ± 0.33	18.16 ± 0.29	15.24 ± 0.27	12.85 ± 0.26	10.51 ± 0.21
significant?	no	no	no	no	yes
p value	1.00	1.00	0.89	0.56	1.7×10^{-2}
ARR					
mRMR	25.19 ± 0.65	20.76 ± 0.63	21.71 ± 0.61	21.64 ± 0.61	-
QPFS $\hat{\alpha} = 0.41$	28.05 ± 0.65	21.30 ± 0.65	21.52 ± 0.65	21.76 ± 0.58	-
significant?	no	no	no	no	-
p value	1.00	0.96	0.69	0.39	-
NCI60					
mRMR	53.50 ± 2.17	34.33 ± 1.74	32.00 ± 1.93	32.83 ± 1.84	33.64 ± 1.80
QPFS $\hat{\alpha} = 0.22$	46.33 ± 2.19	29.83 ± 1.68	29.00 ± 1.83	34.67 ± 1.81	35.17 ± 1.95
significant?	yes	yes	yes	no	no
p value	1.6×10^{-3}	7.5×10^{-3}	3.3×10^{-2}	0.95	0.96
SRBCT					
mRMR	9.38 ± 1.06	2.31 ± 0.51	0.47 ± 0.23	0.24 ± 0.17	0.49 ± 0.30
QPFS $\hat{\alpha} = 0.18$	3.89 ± 0.75	0.11 ± 0.11	0.05 ± 0.11	0.11 ± 0.11	0.35 ± 0.25
significant?	yes	yes	yes	no	no
p value	5.4×10^{-9}	5.6×10^{-5}	2.3×10^{-2}	0.27	0.36
GCM					
mRMR	54.26 ± 1.19	43.38 ± 1.18	41.38 ± 1.08	38.26 ± 1.06	38.50 ± 1.10
QPFS $\hat{\alpha} = 0.15$	65.66 ± 1.03	44.11 ± 1.11	39.57 ± 1.24	38.06 ± 1.16	35.23 ± 1.17
significant?	no	no	yes	no	yes
p value	1.00	0.81	0.037	0.40	1.42×10^{-4}

Table 4: Average error rates using the mRMR and QPFS methods, for classifiers based on M features. The parameter $\hat{\alpha}$ of the QPFS method is indicated; rows are ordered according to this value. The Nyström approximation was used for the GCM and RAT data sets.

and the average number of features selected by Streamwise Feature Selection. SFS was applied to the binary data sets ARR and RAT and was used only as a feature selection method (a feature generation step was not included).

Table 5 shows that for ARR, NCI60, SRBCT and GCM data sets, the best selector is mRMR or QPFS. A statistical study of the performance of both methods is given in Table 4. In the case of the RAT data set, the best methods are MaxRel and reliefF. The fact that the best results are obtained with methods which only consider relevance with the target class fits in with the analysis of Figure 7. Finally, for the MNIST data set the best choice is the mRMR (MIQ) algorithm. Nevertheless, the

performance of MIQ in some data sets is not competitive (see, for instance, the ARR and NCI60 results). The accuracy of QPFS+Nyström ($p = 0.2$) is good if a high enough number of features is used, and it has lower computational cost than mRMR and QPFS.

Regarding SFS, Table 6 shows that SFS provides a competitive error rate for the ARR data set with few features (around 11) but its effectiveness in the RAT data set is improved by other feature selection algorithms when more than 6 attributes are considered. It is noticeable the efficiency of SFS getting acceptable accuracies using a small number of features.

ReliefF and SFS are feature selection methods which need to establish the value of some parameters like in QPFS. In ReliefF all instances were used (not random subsampling) and the number of neighbors was set to 3 for all data sets, except for MNIST where 10 neighbors were considered. In the case of the SFS algorithm, the default values (wealth $= 0.5$ and $\triangle\ \alpha = 0.5$) were used.

3.3 Time Complexity Results

Since the previous subsection has established the effectiveness of the QPFS method, it is useful now to compare mRMR and QPFS experimentally with respect to time complexity. As stated in Table 1 in Section 2.4, the running times of mRMR and QPFS with and without Nyström all depend linearly on N when M and p are fixed. In order to confirm experimentally this theoretical dependence, time consumption as a function of the number of training examples is measured on the SRBCT data set.

Figure 10a shows the time consumed for the modified SRBCT data set, averaged over 50 runs, as a function of the number of samples, N, for the mRMR, QPFS and QPFS+Nyström methods.

As expected, both mRMR and QPFS show a linear dependence on the number of patterns. For QPFS+Nyström, Table 1 shows that the slope of this linear dependence is proportional to the sampling rate p. Over the range $p = 0.01$ to $p = 0.5$, a decrease in p leads to a decrease in the slope of the linear dependence on N. Therefore, although all algorithms are linearly dependent on N, the QPFS+Nyström is computationally the most efficient. The time cost advantage increases with increasing number of training examples because the slope is greater for mRMR than for QPFS.

The next question is the impact on performance of the number of features, M. Table 1 shows that mRMR and QPFS have quadratic and cubic dependence on M, respectively. However, the QPFS+Nyström cubic coefficient is proportional to the square of the sampling rate. When small value of p are sufficient, which is the typical case, the cubic terms are not dominant.

These results are illustrated in the experiments shown in Figure 10b. This figure shows the average time cost for the SRBCT data set as a function of the problem dimension, M, for the mRMR, QPFS, and QPFS+Nyström methods. As expected from Table 1, mRMR and QPFS empirically show quadratic and cubic dependence on problem dimension. QPFS+Nyström shows only quadratic dependence on problem dimension, with a decreasing coefficient for decreasing p values. In all cases, a t-test has been used to verify the order of the polynomial that best fits each curve by least-squares fitting (Neter and Wasserman, 1974). Overall, for small Nyström sampling rates, QPFS+Nyström is computationally the most efficient.

Last but not least important, Table 1 shows there should be a quadratic dependence on sampling rate for the QPFS+Nyström algorithm. Figure 10c shows the empirical average time cost for the SRBCT data set as a function of the sampling rate p. As expected, there is quadratic dependence on p and cubic dependence on the problem dimension M.

Data Set	Method	M						
		10	20	40	50	100	200	400
ARR	MaxRel	27.48	24.68	21.70	20.82	**20.31**	21.73	-
	MID	**25.19**	**22.99**	**20.64**	**20.76**	21.71	**21.64**	-
	MIQ	29.79	27.78	23.89	23.32	21.53	21.74	-
	reliefF	30.64	24.48	21.54	21.34	20.90	21.66	-
	QPFS	28.05	23.72	22.39	21.30	21.52	21.76	-
NCI60	MaxRel	61.33	49.83	40.00	38.67	34.83	35.50	34.17
	MID	53.50	41.50	36.33	34.33	32.00	32.83	**33.67**
	MIQ	56.50	47.50	38.83	38.17	32.83	35.50	35.17
	reliefF	56.93	54.17	48.49	48.49	38.07	**32.13**	34.36
	QPFS	**46.33**	**36.00**	**33.00**	**29.83**	**29.00**	34.67	35.17
SRBCT	MaxRel	21.58	14.33	6.36	4.51	2.19	0.24	**0.13**
	MID	9.39	3.33	2.01	2.31	0.47	0.24	0.49
	MIQ	10.11	2.18	**0.47**	0.72	0.24	0.25	0.72
	reliefF	6.38	4.18	1.65	1.79	0.96	0.40	0.40
	QPFS	**3.89**	**1.57**	0.97	**0.11**	**0.05**	**0.11**	0.35
GCM	MaxRel	79.32	60.78	48.46	45.58	40.98	39.98	38.77
	MID	**54.26**	**48.45**	**44.16**	**43.38**	41.38	38.26	35.50
	MIQ	79.32	56.48	46.64	43.96	41.80	38.46	38.05
	reliefF	61.25	51.61	46.36	43.83	**39.35**	39.75	37.08
	QPFS+N $p = 0.05$	65.66	54.72	46.09	44.11	39.57	**38.06**	**35.26**
RAT	MaxRel	**19.95**	**17.32**	**15.40**	**15.16**	14.34	13.54	11.97
	MID	21.15	18.46	16.53	16.18	14.88	12.81	10.95
	MIQ	23.69	19.62	17.23	16.61	15.07	12.46	10.96
	reliefF	22.16	20.40	17.44	16.45	**13.68**	**11.43**	**9.85**
	QPFS+N $p = 0.1$	27.13	21.89	19.02	18.16	15.24	12.85	10.51
MNIST	MaxRel	59.19	40.98	25.77	22.5	12.09	7.64	6.72
	MID	53.39	29.37	19.56	17.40	11.72	7.55	6.66
	MIQ	51.69	**25.98**	**11.79**	**10.87**	**7.78**	**6.90**	**6.33**
	reliefF	**50.91**	40.20	23.81	19.56	12.31	8.47	6.86
	QPFS+N $p = 0.2$	57.00	35.39	23.62	20.48	11.31	7.71	6.54

Table 5: Error rates for different feature selection methods and Linear SVM. The best result in each case is marked in bold. QPFS+N indicates that the Nyström approximation is used in the QPFS method and p represents the subsampling rate in Nyström method. In all cases, the α parameter of QPFS is set to $\hat{\alpha}$.

4. Conclusions

This paper has presented and studied a new feature selection method for multiclass classifier learning problems. The new method, named Quadratic Programming Feature Selection (QPFS), is based

Data Set	Number of Selected Features (average)	Error rate (%)
ARR	10.75 ± 0.155	23.34 ± 0.63
RAT	6.12 ± 0.13	22.87 ± 0.33

Table 6: Streamwise Feature Selection error rates.

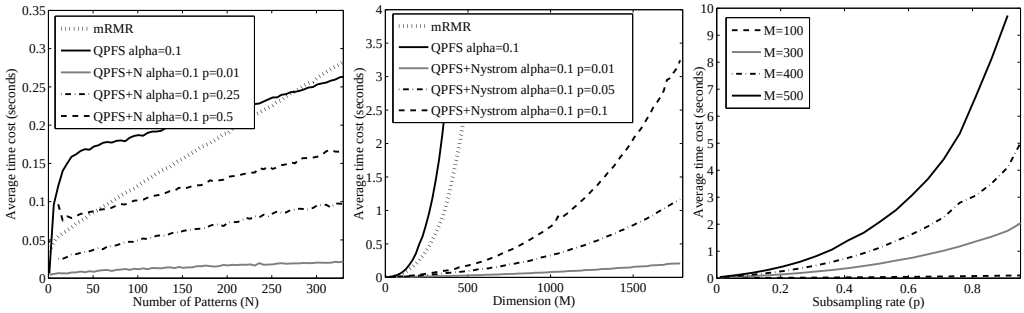

Figure 10: Time cost in seconds for mRMR and QPFS as a function of: **(a)** the number of patterns, N; **(b)** the dimension, M; and **(c)** the sampling rate, p. QPFS+N indicates that the Nyström approximation is used in the QPFS method.

on the optimization of a quadratic function that is reformulated in a lower-dimensional space using the Nyström approximation (QPFS+Nyström). The QPFS+Nyström method, using either Pearson correlation coefficient or mutual information as the underlying similarity measure, is computationally more efficient than the leading previous methods, mRMR and MaxRel.

With respect to classification accuracy, the QPFS method is similar to MaxRel and mRMR when mutual information is used, and yields slightly better results if there is high redundancy. In all experiments, mutual information yields better classification accuracy than correlation, presumably because mutual information better captures nonlinear dependencies. Small sampling rates in the Nyström method still lead to reasonable approximations of exact matrix diagonalization, sharply reducing the time complexity of QPFS. In summary, the new QPFS+Nyström method for selecting a subset of features is a competitive and efficient filter-type feature selection algorithm for high-dimensional classifier learning problems.

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